

Multi-View Channels

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Joint work with



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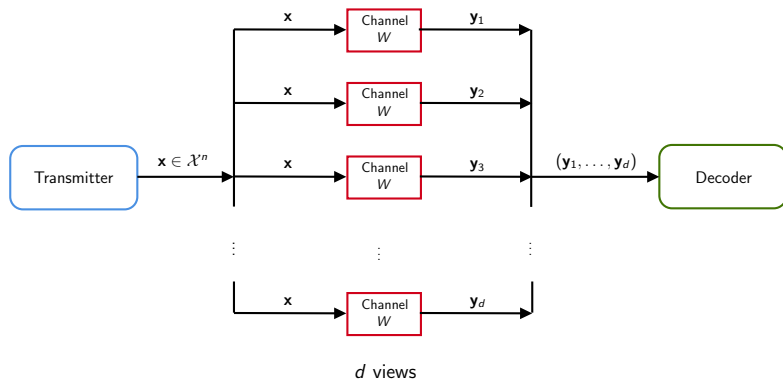


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Wachter-Zeh

(TU München, Germany)

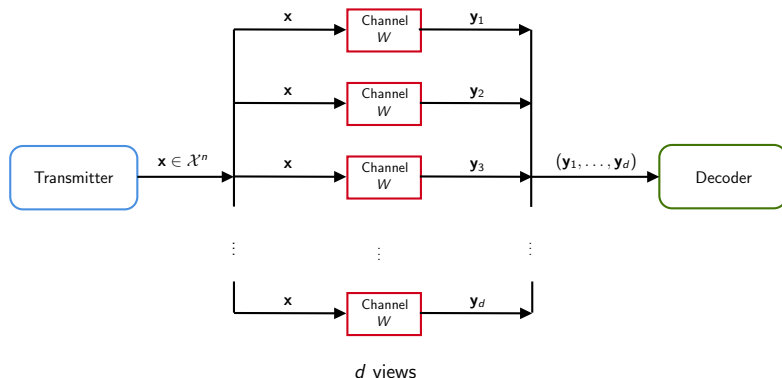
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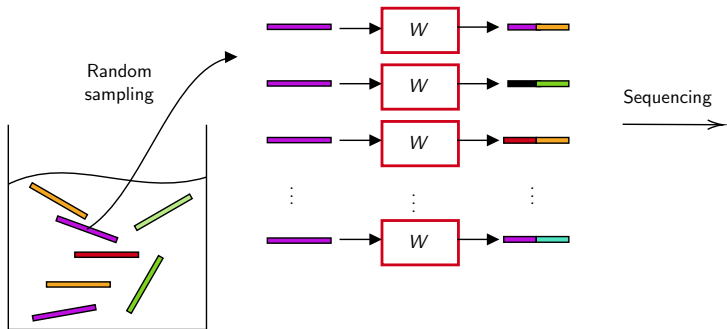


Q1: What rates are achievable, with **vanishing reconstruction error probability**?

Q2: How many views are required on average to **exactly** reconstruct a sequence?

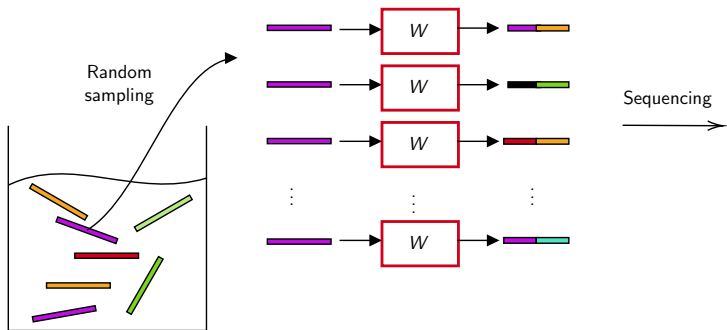
Whence does this problem arise? - I

- ▶ Short molecules in a DNA pool are “amplified” by **Polymerase Chain Reaction** (PCR) and sampled multiple times [Shomorony and Heckel (2022)].



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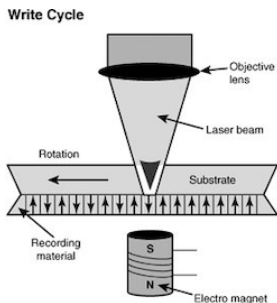
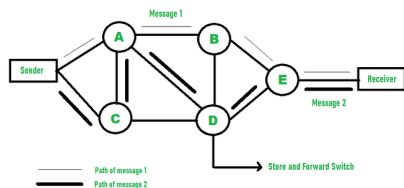
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- ▶ The capacity is determined by mutual information terms corresponding to **different # of views** of the input sequence; see [Lenz et al. (2019, 2020), Shomorony and Heckel (2022), Weinberger and Merhav (2022)].

Whence does this problem arise? - II

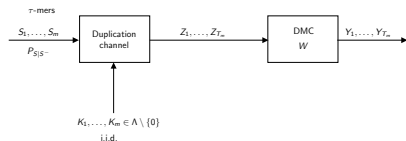
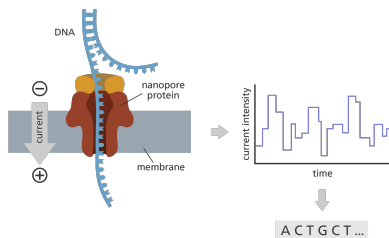
- ▶ Errors due to synchronization in packet-switched communications and in reading magneto-optical media



- ▶ Each output “run” at the end of (noisy) duplications corresponds to a multi-view channel [Mitzenmacher (2008), Cheraghchi and Ribeiro (2019)]

Whence does this problem arise? - III

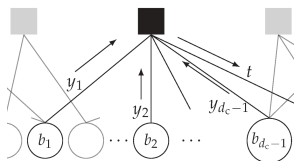
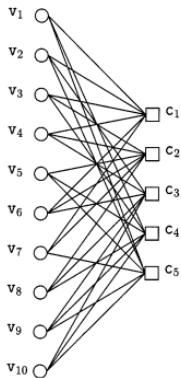
- ▶ Experimentally accurate channel models for DNA nanopore sequencers [McBain and Viterbo (IEEE-BITS 2023), McBain, Viterbo, Saunderson (2024)]



- ▶ The channel model is a noisy duplication channel (with a specific input process).

Whence does this problem arise? - IV

- ▶ More fundamentally, in the iterative decoding of LDPC codes [Gallager (1962), Richardson and Urbanke (2001)]



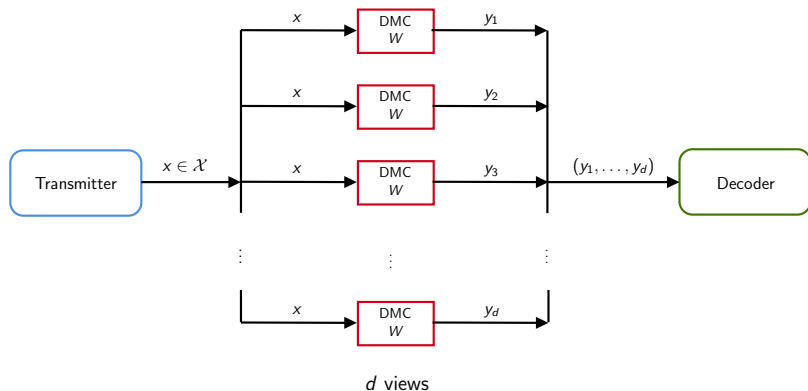
- ▶ A single variable node receives multiple “views” / estimates of its value, from different check nodes.

Answering **Q1** in the large view limit

A simpler setting: Multi-view DMCs

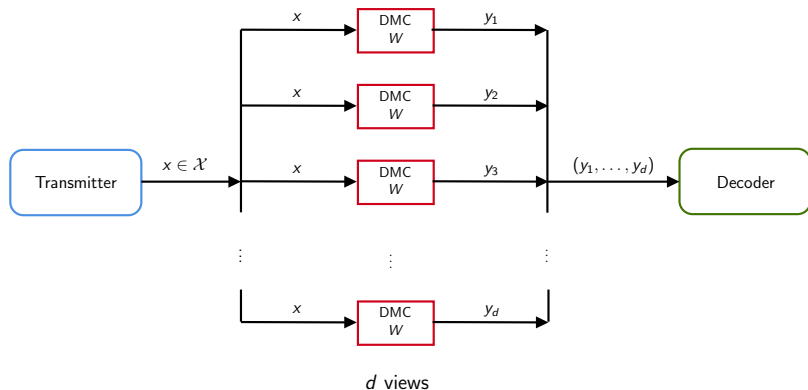
The setting, revisited

The decoder obtains d independent, noisy views of an input symbol.



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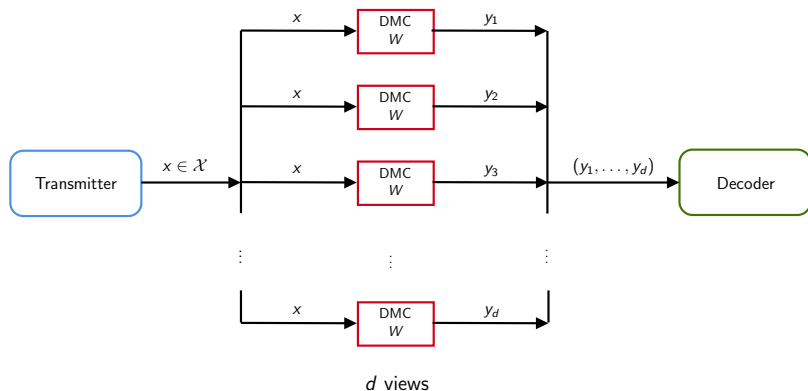
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Since a multi-view DMC is also a DMC, it suffices for us to focus on the transmission of a single symbol, for rate computations.

The setting, revisited

The decoder obtains d independent, noisy views of an input symbol.



Goal: Exact asymptotics of the mutual info. (+ capacity) and dispersion of such a multi-view channel, for arbitrary P_X .

What is known about this setting?

- ▶ Hellman and Raviv (1970), Kanaya and Han (1995): Exact asymptotics of information rates over DMCs with multiple views
- ▶ Levenshtein (2001): Characterization of $\#$ of views for exact reconstruction over comb. error channels and for reconstruction with **decaying** error prob. over multi-view DMCs

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- ▶ Land et al. (2005), Land and Huber (2006): Bounds on the mutual information rates over multi-view DMCs via information combining
- ▶ Mitzenmacher (2006): Calculation of the capacity of a multi-view binary symmetric channel (BSC)
- ▶ Haeupler and Mitzenmacher (2014): Information rates over multi-view deletion channels when $\Pr[\text{deletion}] \rightarrow 0$.
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Our contributions:

1. Unified treatment of info. rate and channel dispersion



Finite-blocklength achievable rates with **fixed** error probability

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Our contributions:

2. Directly extensible proofs for **multi-letter** channels

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Our contributions:

3. **Non-asymptotic capacity bounds** for multi-view BIMS channels

Background

Some formalism

- ▶ Consider a DMC W with input alphabet \mathcal{X} and output alphabet \mathcal{Y} , both finite. Assume that $|\mathcal{X}|, |\mathcal{Y}|$ **do not** depend on d .
- ▶ The d -view **DMC** $W^{(d)}$ obeys the channel law

$$W^{(d)}((y_1, \dots, y_d) | x) = \prod_{i=1}^d W(y_i | x),$$

for $(y_1, \dots, y_d) \in \mathcal{Y}^d$ and $x \in \mathcal{X}$.

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- ▶ Fix a d -independent input distribution P_X . We are interested in:

$$I^{(d)} = I(X; Y^d) = H(X) - H(X | Y^d) \quad \text{[Mutual info.]}$$

and

$$V^{(d)} = \mathbb{E} \left[(\iota(X; Y^d) - I^{(d)})^2 \right], \quad \text{[Channel dispersion]}$$

where

$$\iota(X; Y^d) = \log \frac{P(Y^d | X)}{P(Y^d)}. \quad \text{[Info. spectrum]}$$

Building intuition: Multi-view BSC^(d)

Consider the d -view BSC(p), with binary input $X \in \{0, 1\}$:

$$Y = X + Z \pmod{2},$$

where $Z \sim \text{Ber}(p)$.

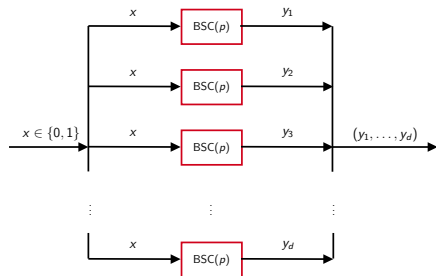
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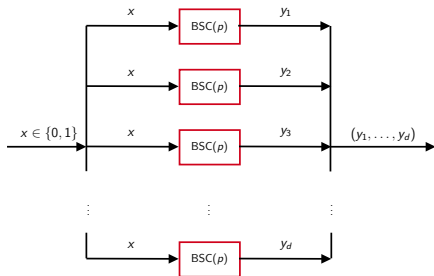
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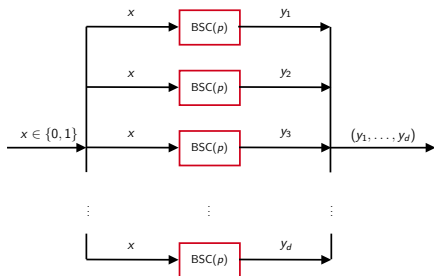
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$$\begin{aligned} I(X; Y^d) &\geq I(X; M) \\ &= \log 2 - h_b(\Pr[X \neq M]). \end{aligned}$$

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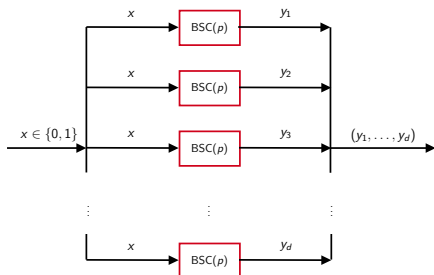
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$$\Pr[X \neq M] \leq \exp(-dZ(p)),$$

where $Z(p) = 2\sqrt{p(1-p)}$.

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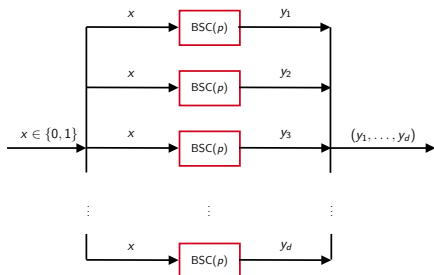
$$\Pr[X \neq M] \leq \exp(-dZ(p)),$$

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Hence, for large d :

$$\begin{aligned} h_b(\Pr[X \neq M]) &\leq -2 \Pr[X \neq M] \cdot \log \Pr[X \neq M] \\ &= 2dZ(p) \cdot \exp(-dZ(p)). \end{aligned}$$

Building intuition: Multi-view BSC^(d)



This gives us

$$I(X; Y^d) \geq \log 2 - 2dZ(p) \cdot \exp(-dZ(p)).$$

Exponentially fast convergence to $H(X) = \log 2$

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Our Results

Our main result

- ▶ Intuitively, as d becomes large, we expect $I^{(d)} \approx H(X)$ and

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- ▶ For distributions P, Q on \mathcal{X} , define the **Chernoff information**

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Theorem

When $\mathcal{X}, \mathcal{Y}, P_X$ do not depend on d ,

$$I^{(d)} = H(X) - \exp(-d\rho + \Theta(\log d|\mathcal{X}|)), \text{ and}$$

$$\left| V^{(d)} - V(X) \right| = \exp(-d\rho + \Theta(\log d|\mathcal{X}|)),$$

where

$$\rho = \min_{x, x': x \neq x'} C(P_{Y|x}, P_{Y|x'}).$$

Interpreting the result

- ▶ The rate of convergence ρ of the mutual information and channel dispersion to their limits

$$\rho = \min_{x, x': x \neq x'} C(P_{Y|x}, P_{Y|x'})$$

is **independent** of the input distribution (except via its support)!

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- ▶ For a binary-input memoryless symmetric (BIMS) channel W , the rate

$$\begin{aligned}\rho &= -\log \sum_{y \in \mathcal{Y}} \sqrt{P_{Y|0}(y|0)P_{Y|1}(y|1)} \\ &= -\log Z_b(W),\end{aligned}$$

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Hence, our earlier speed of convergence for the BSC^(d) is **tight**!

A finite blocklength corollary

- ▶ A characterization of finite-blocklength rates achievable over $W^{(d)}$, thus follows.
- ▶ For a **fixed** $\epsilon \in (0, 1)$ and blocklength $n \geq 1$, let
$$M^*(n, \epsilon) \leftarrow \text{largest } M \text{ s.t. } \exists \text{ length-}n \text{ code over } W^{(d)} \\ \text{with max. error } \epsilon.$$

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Theorem

If W is “non-singular” (*think unlike a BEC*), we have

$$\log |\mathcal{X}| - \frac{\log M^*(n, \epsilon)}{n} \leq e^{-d\rho + \Theta(\log d|\mathcal{X}|)} - \Phi^{-1}(\epsilon) \cdot \frac{e^{-d\rho/2 + \Theta(\log d|\mathcal{X}|)}}{\sqrt{n}} + \Theta\left(\frac{\log n}{n}\right).$$

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- ▶ In particular, choosing $d = \rho^{-1} \log n$, we can achieve rates

$$R_{n, \epsilon} \geq \log |\mathcal{X}| - O_\epsilon\left(\frac{\log n}{n}\right).$$

Proof Sketch I: BIMS channels
bin.-i/p, memoryless, symmetric

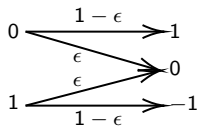
Background on BIMS channels

A BIMS channel obeys:

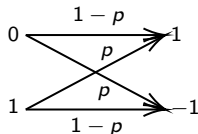
$$Y = (-1)^X \cdot Z,$$

for noise Z independent of X .

Examples:



Binary Erasing Channel (BEC)



Binary Symmetric Channel (BSC)

Proof for BIMS channels with uniform inputs

Consider $I^{(d)}$ for BIMS channels W with $P_X = \text{Ber}(1/2)$.

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Lower bound: $H(X | Y^d) \geq e^{-d\rho + \Theta(\log d)}$

We have

$$H(X | Y^d) = \mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \right].$$

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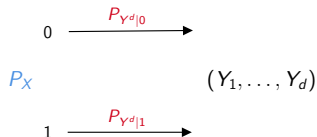
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Note that in the binary hypothesis testing problem



the probability of correct decision is

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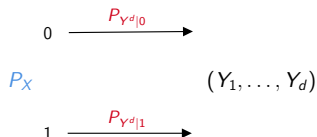
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Applying Jensen's inequality + a little algebra finishes the proof. \square

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A well-known lemma [e.g., [Sasoglu \(2012\)](#)]: for $P_X = \text{Unif}(\{0, 1\})$,

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Here, $W^{(d)}$ is also a BIMS channel, with

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- ▶ [[Hellman & Raviv \(1970\)](#), [Kanaya & Han \(1995\)](#)] (or, our lower bound+[\[Levenshtein \(2001\)\]](#) and [\[Shannon, Gallager, Berlekamp \(1967\)\]](#)) allows us to complete a proof for general DMCs W .
- ▶ Our approach: a **unified proof for $I^{(d)}$ and $V^{(d)}$** that allows for finite-blocklength results+multi-letter extensions.

Proof Sketch II: General DMCs

Sketch of proof strategy for general DMCs

Consider the mutual information $I^{(d)}$ for fixed P_X .

► We write

$$\begin{aligned} H(X | Y^d) &= \mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \right] \\ &= \sum_{x \in \mathcal{X}} P_X(x) \mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \middle| X = x \right]. \end{aligned}$$

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► Fix an $x \in \mathcal{X}$ and focus on the inner term:

$$\mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \middle| X = x \right] = \int_0^\infty \Pr \left[\log \frac{1}{P(x | Y^d)} \geq t \middle| X = x \right] dt.$$

Sketch of proof strategy for general DMCs

Consider the mutual information $I^{(d)}$ for fixed P_X .

► We write

$$\begin{aligned} H(X | Y^d) &= \mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \right] \\ &= \sum_{x \in \mathcal{X}} P_X(x) \mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \middle| X = x \right]. \end{aligned}$$

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Let $p_x(t) := \Pr \left[-\log P_{X|Y^d}(x | Y^d) \geq t \middle| X = x \right]$.

Bounding $p_x(t)$

Proposition

We have that

$$\frac{p_x(t)}{(|\mathcal{X}| - 1)} \leq \max_{\tilde{x} \neq x} \Pr \left[\frac{P(x)P(Y^d | x)}{P(\tilde{x})P(Y^d | \tilde{x})} \leq \frac{e^{-(t - \log(|\mathcal{X}| - 1))}}{1 - e^{-t}} \mid X = x \right]$$

and

$$p_x(t) \geq \max_{\tilde{x} \neq x} \Pr \left[\frac{P(x)P(Y^d | x)}{P(\tilde{x})P(Y^d | \tilde{x})} \leq \frac{e^{-t}}{1 - e^{-t}} \mid X = x \right].$$

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Hence, let us concentrate on terms of the form

$$\Gamma_{x, \tilde{x}}(z) := \Pr \left[\frac{P(x)P(Y^d | x)}{P(\tilde{x})P(Y^d | \tilde{x})} \leq \frac{ce^{-z}}{1 - e^{-z}} \middle| X = x \right],$$

for $c > 0$.

Controlling the integral

Recall that we are primarily interested in

$$\rho_x(t) \approx e^{\Theta(\log |\mathcal{X}|)} \cdot \max_{\tilde{x} \neq x} \Gamma_{x, \tilde{x}}(z),$$

for suitable values of c, z , via $\int_0^\infty \rho_x(t) dt$.

The following observation then holds:

Lemma

We have that

$$\int_0^\infty \max_{\tilde{x} \neq x} \Gamma_{x, \tilde{x}}(z) dz = e^{\Theta(\log |\mathcal{X}|)} \cdot \max_{\tilde{x} \neq x} \int_0^\infty \Gamma_{x, \tilde{x}}(z) dz.$$

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Via **Sanov's theorem**+**Laplace's method**, it is possible to show that

$$\int_0^\infty \Gamma_{x, \tilde{x}}(z) dz = e^{\Theta(\log d|\mathcal{X}|)} \cdot e^{-d \cdot \min_{\tilde{x} \neq x} C(P_{Y|x}, P_{Y|\tilde{x}})}.$$

Putting everything together

Recall that we were ultimately interested in

$$\begin{aligned}\mathbb{E} \left[\log \frac{1}{P(X | Y^d)} \middle| X = x \right] &= \int_0^\infty \Pr \left[\log \frac{1}{P(x | Y^d)} \geq t \middle| X = x \right] dt \\ &= \int_0^\infty p_x(t) dt \\ &\approx e^{\Theta(\log |\mathcal{X}|)} \cdot \int_0^\infty \max_{\tilde{x} \neq x} \Gamma_{x, \tilde{x}}(z) dz \\ &= e^{\Theta(\log d |\mathcal{X}|)} \cdot e^{-d \cdot \min_{\tilde{x} \neq x} C(P_{Y|x}, P_{Y|\tilde{x}})}.\end{aligned}$$

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Hence, we have shown that

$$H(X | Y^d) \approx e^{\Theta(\log d |\mathcal{X}|)} \cdot e^{-d \cdot \min_{\tilde{x} \neq x} C(P_{Y|x}, P_{Y|\tilde{x}})},$$

with $\rho = \min_{\tilde{x} \neq x} C(P_{Y|x}, P_{Y|\tilde{x}})$.

Proof sktech for channel dispersion

Consider the channel dispersion $V^{(d)}$ for fixed P_X .

► We write

$$\begin{aligned} V^{(d)} &= \mathbb{E} \left[(\iota(X; Y^d) - I^{(d)})^2 \right] \\ &= V(X) + \left(\mathbb{E} \left[(\log P(X | Y^d))^2 \right] - H(X | Y^d)^2 \right) + \theta_d, \end{aligned}$$

where

$$\theta_d = 2 \cdot \mathbb{E} \left[(\log P(X | Y^d) + H(X | Y^d)) \cdot \left(\log \frac{1}{P(X)} - H(X) \right) \right].$$

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- ▶ Using an **integral-based approach** (as earlier), we can obtain that

$$\mathbb{E} \left[(\log P(X | Y^d))^2 \right] = e^{-d\rho + \Theta(\log d |\mathcal{X}|)}.$$

- ▶ From the earlier proof, we know that $H(X | Y^d) = e^{-d\rho + \Theta(\log d |\mathcal{X}|)}$.
- ▶ The cross-term θ_d can be handled by “averaging” over $x \in \mathcal{X}$.

Multi-view multi-letter channels

Extensions to multi-letter channels

- ▶ Our main result can be extended to **multi-letter** channels, including synchronization channels **with equal output length across views**.
- ▶ Consider such a multi-letter channel W_n of input length n , with

$$W_n^{(d)}(\mathbf{y}_1, \dots, \mathbf{y}_d \mid x^n) = \prod_{i=1}^d W_n(\mathbf{y}_i \mid x^n)$$

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- ▶ Multi-letter variants of the mutual information $I_n^{(d)}$ and dispersion $V_n^{(d)}$ can be defined similar to earlier.

Theorem

We have that

$$I_n^{(d)} = H(X^n) - \exp(-d\rho_n + \Theta(n \log d |\mathcal{X}|)), \text{ and}$$
$$\left| V_n^{(d)} - V(X^n) \right| = \exp(-d\rho_n + \Theta(n \log d |\mathcal{X}|)),$$

where

$$\rho_n = \min_{u^n, \tilde{u}^n: u^n \neq \tilde{u}^n} C(P_{\mathbf{Y}|u^n}, P_{\mathbf{Y}|\tilde{u}^n}).$$

Computations for the deletion channel

Consider the **deletion channel** $\text{Del}(\delta)$ that, on input $u^n \in \{0, 1\}^n$, deletes each symbol u_i independently w.p. $\delta \in (0, 1)$.

$$0 \mathbf{1} 0 \mathbf{0} \mathbf{1} 0 1 1 \longrightarrow 0 0 0 1 1$$

Theorem

For $\text{Del}(\delta)$, where $\delta \in (0, 1)$, we have that

$$\limsup_{n \rightarrow \infty} \rho_n \leq -\log \delta.$$

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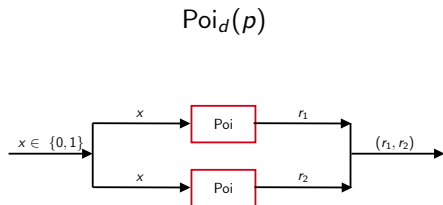
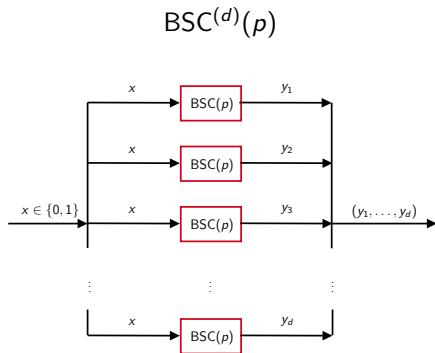
- ▶ The proof follows by obtaining upper bounds on ρ_n by choosing special sequences x^n, \tilde{x}^n .
- ▶ The rate of convergence ρ_n for $\text{Del}(\delta)$, for large enough n , is hence **much slower** than that of an n -letter DMC W_n , for which

$$\rho_n(W_n) = n \cdot \rho(W).$$

Non-Asymptotic Bounds on Capacity

Non-asymptotic bounds for the BSC^(d)

Consider the BSC^(d)(p). We introduce a closely related DMC Poi _{d} (p).

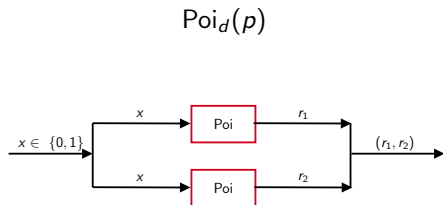
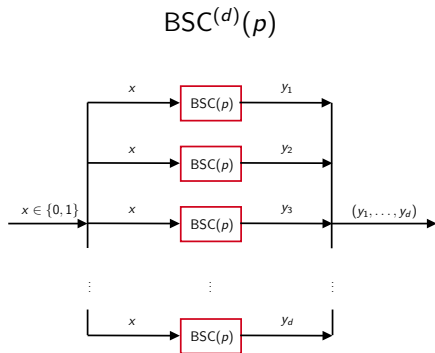


$$P_{R_1|0} = \text{Poi}(d(1-p)), \quad P_{R_1|1} = \text{Poi}(dp)$$

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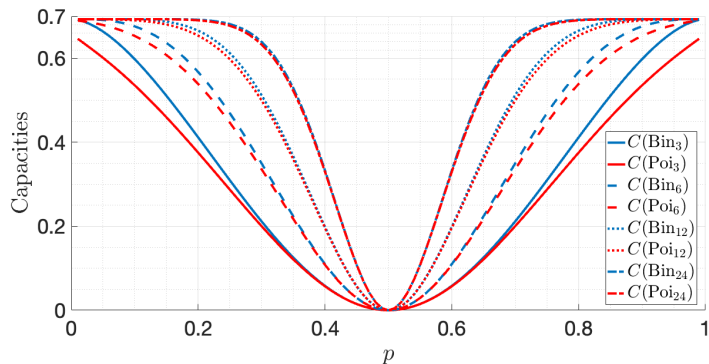
$$P_{R_2|1} = \text{Poi}(d(1-p)), \quad P_{R_2|0} = \text{Poi}(dp)$$

Theorem

$$C(\text{Poi}_d(p)) \leq C(\text{Bin}_d(p)) \leq C(\text{Poi}_d(p)) + e^{-d(1-Z(p))} - Z(p)^{2d},$$

where $Z(p) = 2\sqrt{p(1-p)}$.

Plots and comparisons



Plots comparing $C(\text{Bin}_d(p))$ and $C(\text{Poi}_d(p))$, for varying d .

A corollary for BIMS channels

- ▶ Any BIMS channel W with a finite output alphabet can be decomposed as

$$W_{Y|X=x} = \sum_{i=1}^K \epsilon_i \cdot W_{Y|X=x}^{(i)},$$

where $W_{Y|X=x}^{(i)}$ is the channel law of a BSC with crossover probability p_i , and $\epsilon_i > 0$, with $\sum_i \epsilon_i = 1$.

- ▶ Let P_s be the distribution on $\{1, \dots, K\}$ with mass ϵ_i at point i .

$$C^{(d)}(W) = \mathbb{E}[C(J_1, \dots, J_d)],$$

where $J_i \stackrel{\text{i.i.d.}}{\sim} P_s$, $1 \leq i \leq K$.

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- Now, suppose that $0 < p_1 \leq p_2 \leq \dots \leq p_K < 1/2$. Then,

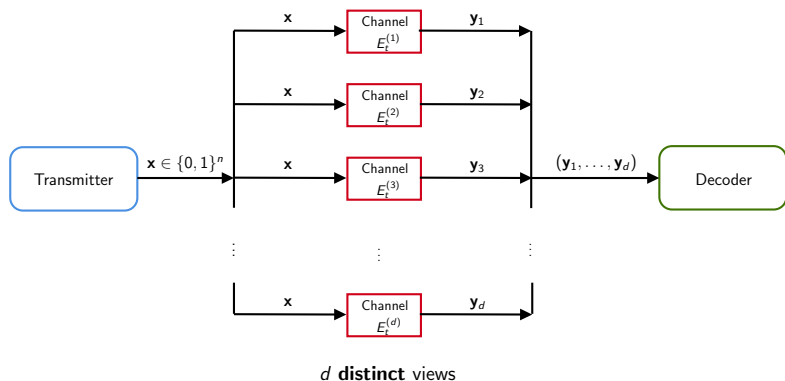
Corollary

$$\begin{aligned} C(\text{Poi}_d(p_K)) &\leq C^{(d)}(W) \\ &\leq C(\text{Poi}_d(p_1)) + \exp(-d(1 - Z(p_1))) - Z(p_1)^{2d}. \end{aligned}$$

Towards **Q2**: Extending Levenshtein's results to the average case

The setting

Consider the following special setting:



- ▶ The views $\mathbf{y}_1, \dots, \mathbf{y}_d$ are **distinct** and are drawn uniformly, without replacement, from a t -substitution error sphere around \mathbf{x} , i.e.,

$$\mathbf{y}_i = \mathbf{x} + \mathbf{e}_i, \text{ where } w_H(\mathbf{e}_i) = t \text{ and } \mathbf{e}_1, \dots, \mathbf{e}_d \text{ are distinct.}$$

Some comments

- ▶ This setting is a generalization of the classical problem in [Levenshtein (2001)], wherein the (distinct) errors $\mathbf{e}_1, \dots, \mathbf{e}_d$ were adversarially drawn from the Hamming sphere of radius t .

Some comments

- ▶ This setting is a generalization of the classical problem in [Levenshtein (2001)], wherein the (distinct) errors $\mathbf{e}_1, \dots, \mathbf{e}_d$ were adversarially drawn from the Hamming sphere of radius t .
- ▶ From [Levenshtein (2001)], it can be argued that

$$\text{Dec}(\mathbf{y}_1, \dots, \mathbf{y}_d) = \mathbf{x},$$

for all $\mathbf{x} \in \{0, 1\}^n$, only if $d > N_{\text{Lev}} := 2 \binom{n-1}{t-1}$, for any decoder Dec.

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- ▶ Moreover, Levenshtein argued that the bit-wise majority decoder

$$\text{Dec}^{\text{Maj}}(\mathbf{y}_1, \dots, \mathbf{y}_d) = (\text{Maj}(y_{1,1}, \dots, y_{d,1}), \dots, \text{Maj}(y_{1,n}, \dots, y_{d,n}))$$

is “optimal”, in that

$$\text{Dec}^{\text{Maj}}(\mathbf{y}_1, \dots, \mathbf{y}_d) = \mathbf{x},$$

for all $\mathbf{x} \in \{0, 1\}^n$, whenever $d > N_{\text{Lev}}$.

How does Dec^{Maj} perform in the average case?

Fix the bit-wise majority decoder Dec^{Maj} . Given views $\mathbf{y}_1, \dots, \mathbf{y}_d$, construct the array

$$M(\mathbf{y}_1, \dots, \mathbf{y}_d) =$$

$y_{1,1}$	$y_{1,2}$					$y_{1,n}$	\mathbf{y}_1
$y_{2,1}$							\mathbf{y}_2
							\mathbf{y}_3
							\vdots
$y_{d,1}$						$y_{d,n}$	\mathbf{y}_d

Let $m(n, t, d)$ denote the number of distinct arrays $M(\mathbf{y}_1, \dots, \mathbf{y}_d)$ that lead to correct reconstruction of \mathbf{x} via Dec^{Maj} .

A recurrence relation for estimating $m(n, t, d)$

From the structure of the Dec^{Maj} decoder, the following lemma holds:

Lemma

We have that

$$m(n, t, d) \geq m(n-3, t, d) + m(n-3, t-1, d) \cdot a_d,$$

where $a_d = \sum_{\substack{k_1+k_2+k_3=N \\ k_1, k_2, k_3 \leq \lfloor \frac{d-1}{2} \rfloor}} \binom{d}{k_1, k_2, k_3}.$

- ▶ With the aid of suitable initial conditions

$$m(n, t, d) = 0, \text{ if } d < 3 \text{ or } d > n - 3(t - 1),$$

the above recurrence relation can be explicitly solved, giving a **lower bound** on $m(n, t, d)$.

Estimating the average # of views for reconstruction

- ▶ Via the previous arguments, we can derive a lower bound on

$$P_c(n, t, d) := \frac{m(n, t, d)}{\binom{n}{d}}.$$

- ▶ This allows us to obtain an **upper bound** on $\mathbb{E}[\# \text{ views for reconst.}]$ via a union-bound style argument.
- ▶ Moreover, we have the following lemma:

Lemma

When $t = 1$, we have

$$\mathbb{E}[\# \text{ views for reconst.}] = N_{\text{Lev}} = 3.$$

Future work

- ▶ Bounds/exact computation of ρ_n for **general synchronization channels** and other **channels with memory**
- ▶ Relating Dec^{Maj} to the “**optimal**” decoder in terms of $\mathbb{E}[\# \text{ views for reconst.}]$.

References

1. V. Arvind Rameshwar and Nir Weinberger, “**Information rates over multi-view channels,**” under review at the [IEEE Transactions on Information Theory](#); presented at the [IEEE ISIT 2024](#).
2. Vivian Papadopoulou, V. Arvind Rameshwar, Antonia Wachter-Zeh, “**How many noisy copies of a sequence are needed on average for exact reconstruction?,**” accepted to the [ITW 2024](#).

Thank You!